## ACSC/STAT 3703, Actuarial Models I

WINTER 2025 Toby Kenney Homework Sheet 3

Model Solutions

## **Basic Questions**

1. A distribution has hazard rate  $\lambda(x) = x + \frac{3}{6+x}$  for  $x \ge 0$ . Calculate its density function.

The survival function is

$$S(x) = e^{-\int_0^x \lambda(s) \, ds} = e^{-\int_0^x x + \frac{3}{6+y} \, dy}$$
$$= e^{-\left[\frac{y^2}{2} + 3\log(6+y)\right]_0^x}$$
$$= e^{-\left(\frac{x^2}{2} + 3(\log(6+x) - \log(6))\right)}$$
$$= \frac{6^3}{(6+x)^3} e^{-\frac{x^2}{2}}$$

The density function is negative the derivative of the survival function. That is

$$f(x) = -\frac{dS(x)}{dx}$$
  
=  $-\frac{d}{dx} \left( \frac{6^3}{(6+x)^3} e^{-\frac{x^2}{2}} \right)$   
=  $\frac{3 \times 6^3}{(6+x)^4} e^{-\frac{x^2}{2}} + \frac{6^3 x}{(6+x)^3} e^{-\frac{x^2}{2}}$   
=  $\frac{6^3(3+6x+x^2)}{(6+x)^4} e^{-\frac{x^2}{2}}$ 

2. A continuous random variable has moment generating function given by  $M(t) = \frac{1}{(1-t)^2} + \frac{1}{(1-t^2)^4}$ . What is the skewness of the distribution?

We calculate

$$M'(t) = \frac{2}{(1-t)^3} + \frac{8t}{(1-t^2)^5}$$
$$M''(t) = \frac{6}{(1-t)^4} + \frac{8}{(1-t^2)^5} + \frac{80t^2}{(1-t^2)^6}$$
$$M'''(t) = \frac{24}{(1-t)^5} + \frac{240t}{(1-t^2)^5} + \frac{960t^3}{(1-t^2)^6}$$
$$M'(0) = 2$$
$$M''(0) = 14$$
$$M'''(0) = 24$$

Thus  $\mathbb{E}(X) = 2$  and  $\mathbb{E}(X^2) = 14$  and  $\mathbb{E}(X^3) = 24$ . This gives  $Var(X) = 14 - 2^2 = 10$  and

$$\mu_3 = \mathbb{E}(X^3) - 3\mathbb{E}(X)\mathbb{E}(X^2) + 2\mathbb{E}(X)^3 = 24 - 3 \times 2 \times 14 + 2 \times 2^3 = -44$$
  
Thus, the skewness is  $\frac{-44}{10^{1.5}} = -1.39140217047$ 

3. Calculate the mean excess loss function for a distribution with hazard rate given by  $\lambda(x) = \frac{x}{x+1}$  for  $x \ge 0$ .

Rewriting  $\lambda(x) = 1 - \frac{1}{x+1}$ , we have

$$S(x) = e^{-\int_0^x \lambda(y) \, dy} = e^{-\int_0^x 1 - \frac{1}{y+1} \, dy} = e^{-x + \log(x+1)} = (x+1)e^{-x}$$

The mean excess loss function is given by

$$\mathbb{E}((X-d)_{+}) = \int_{d}^{\infty} S(x) \, dx$$
  
=  $\int_{d}^{\infty} (x+1)e^{-x} \, dx$   
=  $[-(x+1)e^{-x}]_{d}^{\infty} + \int_{d}^{\infty} e^{-x} \, dx$   
=  $(d+1)e^{-d} + e^{-d}$   
=  $(d+2)e^{-d}$ 

4. Calculate the probability generating function of a discrete distribution with p.m.f. given by

$$f(n) = \frac{e^{-1}}{2} \frac{n^2}{n!}$$

The probability generating function is given by

$$P(z) = \mathbb{E}(z^X) = \sum_{n=0}^{\infty} f(x) z^x$$
  
=  $\frac{e^{-1}}{2} \sum_{n=0}^{\infty} \frac{n^2 z^n}{n!}$   
=  $\frac{e^{-1}}{2} \sum_{n=1}^{\infty} \frac{n z^n}{(n-1)!}$   
=  $\frac{e^{-1}}{2} \sum_{n=1}^{\infty} \frac{(n-1+1)z^n}{(n-1)!}$   
=  $\frac{e^{-1}}{2} \left( \sum_{n=1}^{\infty} \frac{z^n}{(n-1)!} + \sum_{n=2}^{\infty} \frac{z^n}{(n-2)!} \right)$   
=  $\frac{e^{-1}}{2} \left( z \sum_{m=0}^{\infty} \frac{z^m}{m!} + z^2 \sum_{m=0}^{\infty} \frac{z^m}{m!} \right)$   
=  $\frac{z+z^2}{2} e^{z-1}$ 

## **Standard Questions**

5. The total cost of handling a claim is  $X_1 + X_2 + Y$  where  $X_1$  and  $X_2$  are i.i.d. discrete non-negative random variables with probability generating function  $P_X(z) = \frac{z+1}{z+3}$  and Y is a discrete non-negative random variable with probability generating function  $P_Y(z) = \frac{3^z}{4(z+1)}$ , independent of  $X_1$ and  $X_2$ . What is the moment generating function of  $X_1 + X_2 + Y$ ?

We have

$$P_{X_1+X_2+Y}(z) = P_{X_1}(z)P_{X_2}(z)P_Y(z)$$
  
=  $P_X(z)^2 P_Y(z)$   
=  $\left(\frac{z+1}{z+3}\right)^2 \frac{3^z}{4(z+1)}$   
=  $\frac{(z+1)3^z}{4(z+3)^2}$ 

The moment generating function is given by  $M_X(t) = P_X(e^t)$ , so

$$M_{X_1+X_2+Y}(t) = P_{X_1+X_2+Y}(e^t) = \frac{(e^t+1)3^{e^t}}{4(e^t+3)^2}$$

6. An insurance company is considering two models for its data. The first is a Pareto distribution with survival function

$$S(x) = \left(\frac{\theta}{\theta + x}\right)^{\alpha}$$

The second is a Weibull distribution with survival function

$$S(x) = e^{-\left(\frac{x}{\theta}\right)^{\prime}}$$

They find that for the fitted parameters, both distributions have the same values of  $\theta$ , and the same values for the 90th and 95th percentiles. Which distribution has a higher 99th percentile?

[You should get an equation for one of the unknown parameters  $\alpha$  or  $\tau$ . You can numerically solve this equation by trying a range of values and seeing which satisfies the equation.]

The pth percentiles of the Pareto and Weibull distributions are obtained by solving

$$\left(\frac{\theta}{\theta+x}\right)^{\alpha} = 1-p$$
$$x = \theta(1-p)^{-\alpha} - 1$$
$$e^{-\left(\frac{x}{\theta}\right)^{\tau}} = 1-p$$
$$x = \theta(-\log(1-p))^{\frac{1}{\tau}}$$

Setting the 90th and 95th percentiles equal gives:

$$\begin{aligned} \theta(-\log(0.1))^{\frac{1}{\tau}} &= \theta(0.1^{-\alpha} - 1) \\ \theta(-\log(0.05))^{\frac{1}{\tau}} &= \theta(0.05^{-\alpha} - 1) \\ \frac{\log(-\log(0.1))}{\tau} &= \log\left((0.1)^{-\alpha} - 1\right) \\ \frac{\log(-\log(0.05))}{\tau} &= \log\left((0.05)^{-\alpha} - 1\right) \\ \frac{\log\left((0.1)^{-\alpha} - 1\right)}{\log\left((0.05)^{-\alpha} - 1\right)} &= \frac{\log(-\log(0.1))}{\log(-\log(0.05))} \\ &= 0.760154060074 \\ \log\left((0.1)^{-\alpha} - 1\right) &= 0.760154060074 \log\left((0.05)^{-\alpha} - 1\right) \\ (0.1)^{-\alpha} - 1 &= \left((0.05)^{-\alpha} - 1\right)^{0.760154060074} \end{aligned}$$

Numerically, we see that this is solved by  $\alpha = 1.3228$ . Substituting this

for  $\alpha$ , we get

$$\frac{\log(-\log(0.1))}{\tau} = \log\left((0.1)^{-1.3228} - 1\right)$$
$$\frac{0.834032445246}{\tau} = 2.9971362061$$
$$\tau = \frac{0.834032445246}{2.9971362061}$$
$$\tau = 0.2783$$

The 99th percentiles are

$$\theta\left((0.01)^{-1.3228} - 1\right) = 441.180921074\theta$$

for the Pareto distribution, and

$$\theta(-\log(0.01))^{\frac{1}{0.278276457222}} = 241.771811347\theta$$

for the Weibull distribution, so the percentile is larger for the Pareto distribution.

## **Bonus Questions**

7. X and Y are continuous random variables with moment generating functions  $M_X(t) = \frac{864}{(t-4)(t-6)^3}$  and  $M_Y(t) = \frac{e^{-t^2}}{t+1}$ . You are given that X and X + Y are independent. What is the probability generating function of X + Y?

Since X and X + Y are independent, -X and X + Y are independent. so  $M_Y(t) = M_{-X}(t)M_{X+Y}(t)$ . We have

$$M_{-X}(t) = \mathbb{E}\left(e^{t(-X)}\right) = \mathbb{E}\left(e^{-tX}\right) = M_X(-t) = \frac{864}{(t+4)(t+6)^3}$$

so we have

$$\frac{e^{-t^2}}{t+1} = \frac{864}{(t+4)(t+6)^3} M_{X+Y}(t)$$
$$M_{X+Y}(t) = -\frac{(t+4)(t+6)^3 e^{-t^2}}{864(t+1)}$$

The probability generating function satisfies  $P(z) = M(\log(z))$ , so

$$P_{X+Y}(z) = -\frac{(\log(z) + 4)(\log(z) + 6)^3 e^{-\log(z)^2}}{864(\log(z) + 1)}$$